Demo 28: The Confidence Interval of a Proportion

Defining the confidence interval • Looking at sample results in terms of plausibility

This demo is about the meaning of a confidence interval (CI). The CI is a difficult concept; it's easy to get confused. We can talk about how the CI measures confidence in a *process*—the process of constructing the interval—and that the confidence level is the probability that if the process were to be repeated, such-and-such a percentage of the intervals so constructed would contain the true population value. While that's accurate, it sounds circular and is mired in subjunctive, subtle language.

I prefer this definition, which is the basis of our demo:

The confidence interval is the range of possible parameter values for which the observed value is *plausible*.

What do we mean by "plausible"? Let's look at the rest of the statement in context. Suppose we have poll results that say 56% of people support Measure Q. This poll is of a particular number of people N, the sample size. We assume that it's a simple random sample and that N is much smaller than the population.

The population parameter we're interested in is p, the proportion of people who favor Measure Q. (In other contexts, we might be interested in the *mean* height of students or the *median* income of dock workers. But this is the confidence interval of a *proportion*; the logic will be the same.) Suppose we polled 50 people and 28 of them said they favored Q. That gives us the sample proportion 0.56—and is where the 56% came from. We call the sample proportion \hat{p} , pronounced p-hat. We're interested in the population parameter p, but we will never know its value.

Instead, we ask, "Would we be surprised if the true population parameter were 54%?" Certainly not. It would be easy to get 28 positive responses out of 50 if the population had 54% in favor. But would it be plausible for only 10% of the population to be in favor? No. The chances of getting 28 out of 50 are too small. Traditionally, the border between plausible and implausible is a 5% probability, though you can change that depending on your situation. All that's left is to find that range of values for which the population parameter is plausible.

That's a long introduction; let's see what it means:

What To Do

Open Cl of a Proportion.ftm. It will look something like the illustration.

This document simulates many repeated polls. The **Poll** collection (upper left) contains 50 cases, **yes** or **no**, randomly controlled by the **p_test** slider. That is, **p_test** is the probability that a respondent in **Poll** will vote **yes**. The bar graph shows the results of *one* of those polls.

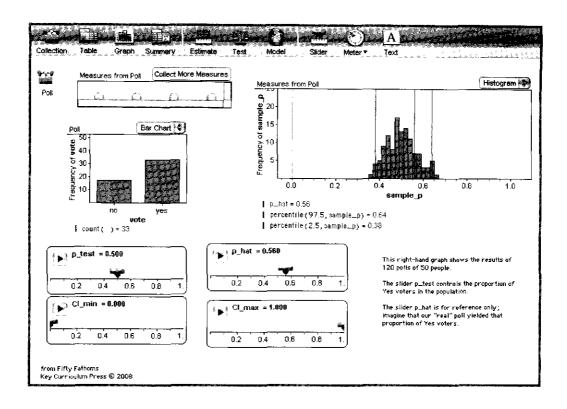
The **Measures from Poll** collection and graph show the results of 120 simulated polls of 50 people. In the graph, the two "outside" vertical lines show the 2.5% and 97.5% percentiles for the polls in the graph; that is,

95% of the polls are between the two lines. We will call all of those results *plausible*.

The **p_hat** slider controls a line on the graph; it represents the results of our real-world poll. So the graph in the illustration, with a population proportion of **p_test = 0.50**, shows that our poll, **p_hat**, is a plausible result.

Note that the histogram of **sample_p**—the simulated poll results—is very much like the graph in Demo 20, "The Distribution of Sample Proportions."

Drag the slider called **p_test**. The graph will slowly update (it has to simulate 120 polls of 50 people every time). Set it to 0.3. For that value, you



can see that our **p_hat**, 0.56, is outside the bounds of plausibility. If 30% of the true population favored Measure Q, it would be really unlikely that we would get a 56% result in a poll of 50 people.

- Experiment with different values of p_test to find values so that the upper and lower "boundary" lines on the graph are coincident with the p_hat line. (In fact, since Fathom draws the p_hat line first, these "percentile" lines will hide the red p_hat line when they match exactly.)
 - ⇒ If you set p_test to unusual values (such as -1), the histogram may look strange. Just choose Histogram again from the menu in the corner of the graph.

Question

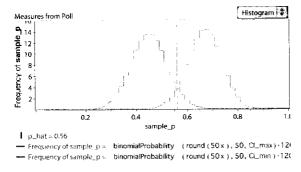
1 Why do you suppose the "percentile" values below the graph seem to be mostly multiples of 2%? That is, you might get 0.56 or 0.58, but seldom 0.57, and never 0.572.

Onward!

- When you find a proportion that seems the minimum p_test for which p_hat is plausible, set the slider Cl_min to that value. When you find the maximum value, set Cl_max.
- Let's draw the theoretical distribution functions for population proportions at those two limits.
 Select the graph and choose Plot Function from the Graph menu. The formula editor opens.
- Enter binomialProbability(round(50x), 50,Cl_max)*120. Press OK to plot the function and close the editor.
 - ⇒ For help with this function, see "Plotting Binomial Probability—and Other Discrete Distributions" in Appendix A.
- Choose Plot Function again; this time enter binomialProbability(round(50x), 50,Cl_min)*120 (that is, the same except for "min" instead of "max"). Press OK to close the editor.
- With the data on the graph, it's pretty busy. Choose Select All from the Edit menu to select all the cases in the graph; then choose Delete Cases from the Edit menu to get rid of them.

We also have those leftover vertical lines. Select the first one's formula at the bottom of the graph (it starts **percentile(97.5,...)** and choose **Clear Formula** from the **Edit** menu. The line evaporates. Do the same to the other **percentile** line.

Now you have only the functions and **p_hat** left. The graph will look like this:



Note: Your **p_hat** line may not cut the curves exactly at the intersection.

Now our display shows us how the distributions of the minimum and maximum of the interval relate to **p_hat**. The **p_hat** line cuts off two tails of the distributions; each has an area of approximately 2.5%. Those tails represent the chance that we are wrong, that our interval fails to enclose the true population proportion.

If you want less risk of making that mistake, move the **CI_min** and **CI_max** sliders to make the tail bits smaller. Then the peaks will be farther apart, demonstrating in a different way why confidence intervals get wider—less precise—if you increase the confidence level.

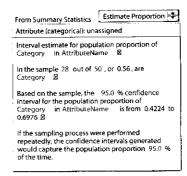
Move p_hat to a different value, then move the Cl_min and Cl_max sliders to cut off about the right amount of tails.

Another Question

What happens when you set **p_hat** to a very large or small number, like 0.96 or 0.02? What values of **Cl_max** and **Cl_min** are appropriate?

Extensions

- ➤ Test your interval by using a traditional confidence interval: drag a new estimate from the shelf. (It looks like a ruler.)
- ➤ Choose Estimate Proportion from the pop-up menu in the estimate itself.



- Enter the data directly into the estimate (do not drag an attribute there). For example, make it look like the illustration.
- Compare your "plausibility interval" values to the ones you see in the estimate.
- Change the confidence level to 90% and redo both the plausibility interval (as above) and the one in the estimate.

Challenges

- 3 Explain clearly why this "plausibility interval" is really the same as the traditional confidence interval ("if you were to draw new samples and construct intervals repeatedly . . .") described at the beginning of this demo.
- 4 Explain why, to get a 95% confidence interval, we use the 2.5 and 97.5 percentiles. That makes sense if you have one distribution (the two tails together make 5%), but here we have two distributions with a combined area of 2, not 1.
- What if **p_hat** is zero or one? Can you still make a "plausibility interval"? Explain what it means, and relate it to the traditional confidence interval and what you get using Fathom's estimate.